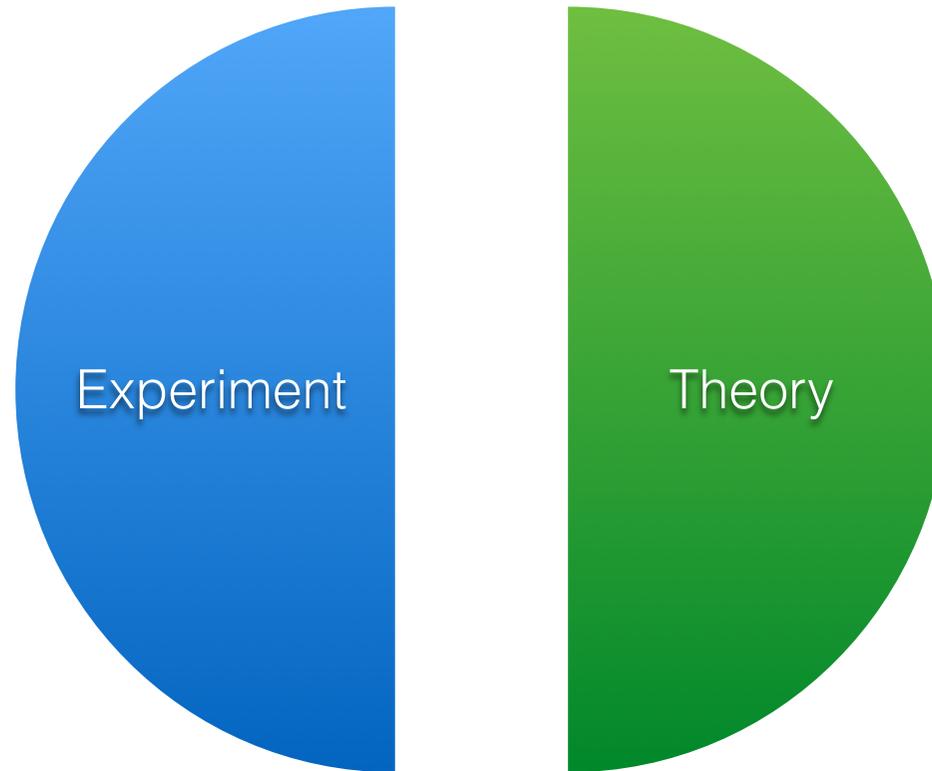


Statistics for Astronomy

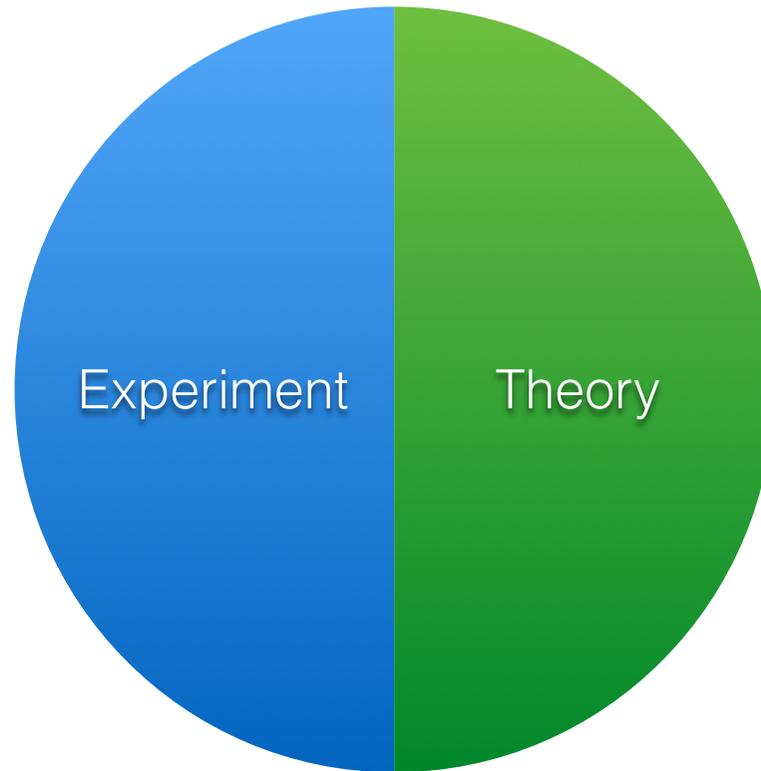
Statistical Inference

Adam Amara
ETH zürich

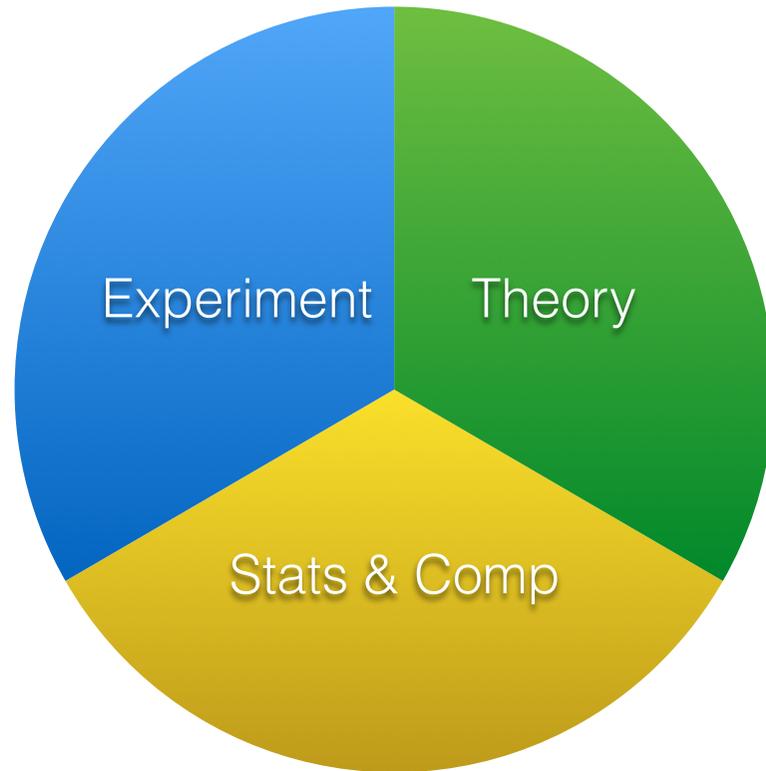
Domains in Physics



Domains in Astrophysics



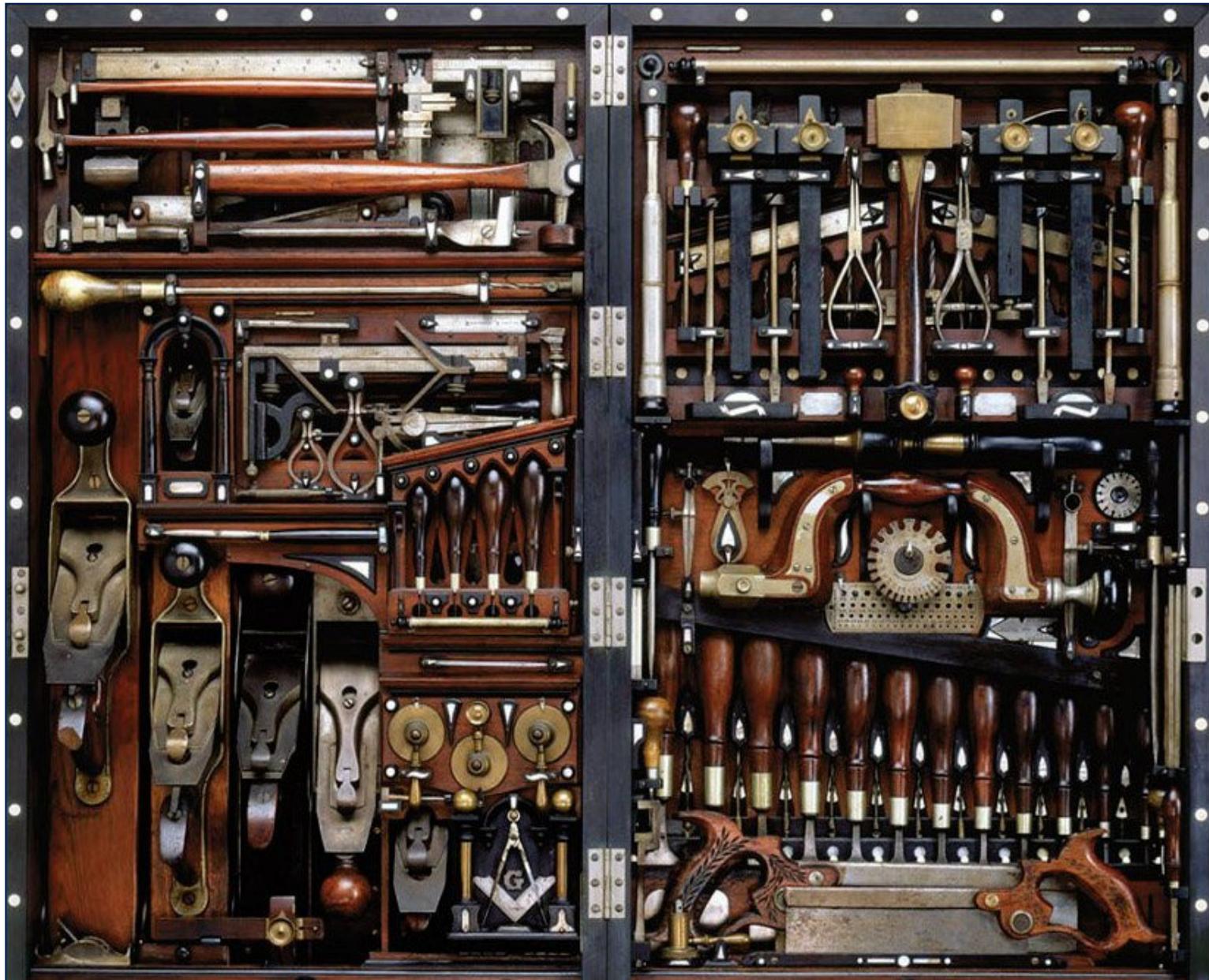
Domains in Astrophysics



The Aim



The Aim



Bayesians, Frequentists and other

Probability

$$p(\theta|d) \propto p(d|\theta)p(\theta)$$

‘prior’

‘degree of belief’

DID THE SUN JUST EXPLODE?
(IT'S NIGHT, SO WE'RE NOT SURE.)

THIS NEUTRINO DETECTOR MEASURES
WHETHER THE SUN HAS GONE NOVA.

THEN, IT ROLLS TWO DICE. IF THEY
BOTH COME UP SIX, IT LIES TO US.
OTHERWISE, IT TELLS THE TRUTH.

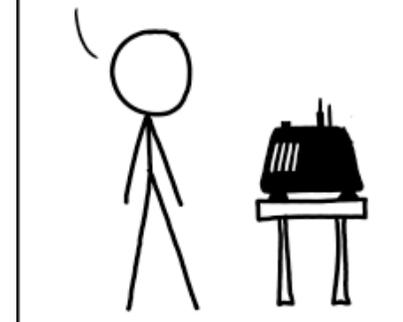
LET'S TRY.

DETECTOR! HAS THE
SUN GONE NOVA?



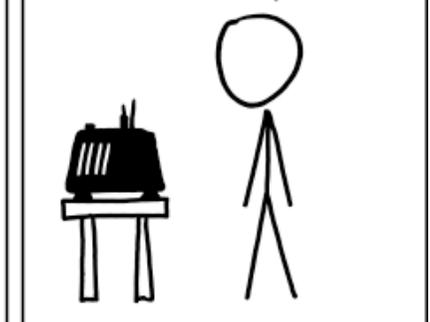
FREQUENTIST STATISTICIAN:

THE PROBABILITY OF THIS RESULT
HAPPENING BY CHANCE IS $\frac{1}{36} = 0.027$.
SINCE $p < 0.05$, I CONCLUDE
THAT THE SUN HAS EXPLODED.



BAYESIAN STATISTICIAN:

BET YOU \$50
IT HASN'T.



Experimental errors and uncertainties
should be treated probabilistically

Treatment of probabilities should be
mathematically sound

Interpretation of probability (subjective vs
objective) depends on the situation

Bayesians, Frequentists and other

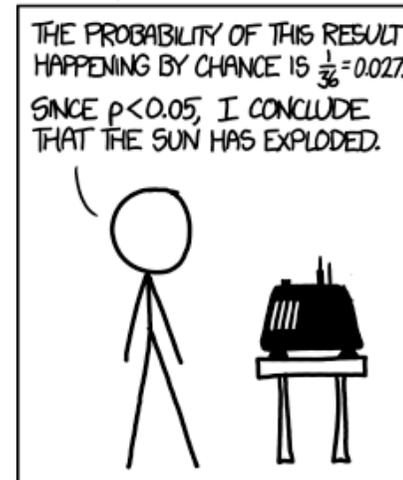
It is possible to do garbage statistics as a 'Bayesian' or 'Frequentist': i.e. just because you use the word 'prior', does not mean you know what you're doing

Best is to understand the structure of your problem (e.g. your error model) and to use mathematically sound probabilities.

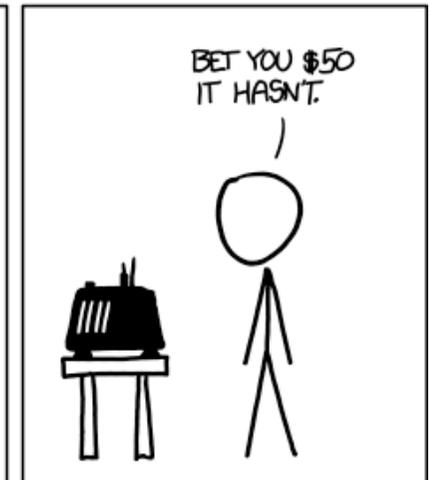
DID THE SUN JUST EXPLODE?
(IT'S NIGHT, SO WE'RE NOT SURE.)



FREQUENTIST STATISTICIAN:



BAYESIAN STATISTICIAN:

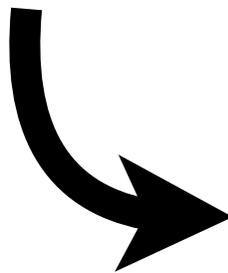




Good overview of available tools

Well organised and maintained

Know which tool for which job



Relative Entropy

Hierarchical Models

K-S test

Maximum Likelihood

p-value

Joint Probability

Minimum ' χ^2 '

Hypothesis testing

Graphical Models

Bayes Theorem

MCMC

Host of Classical Tests

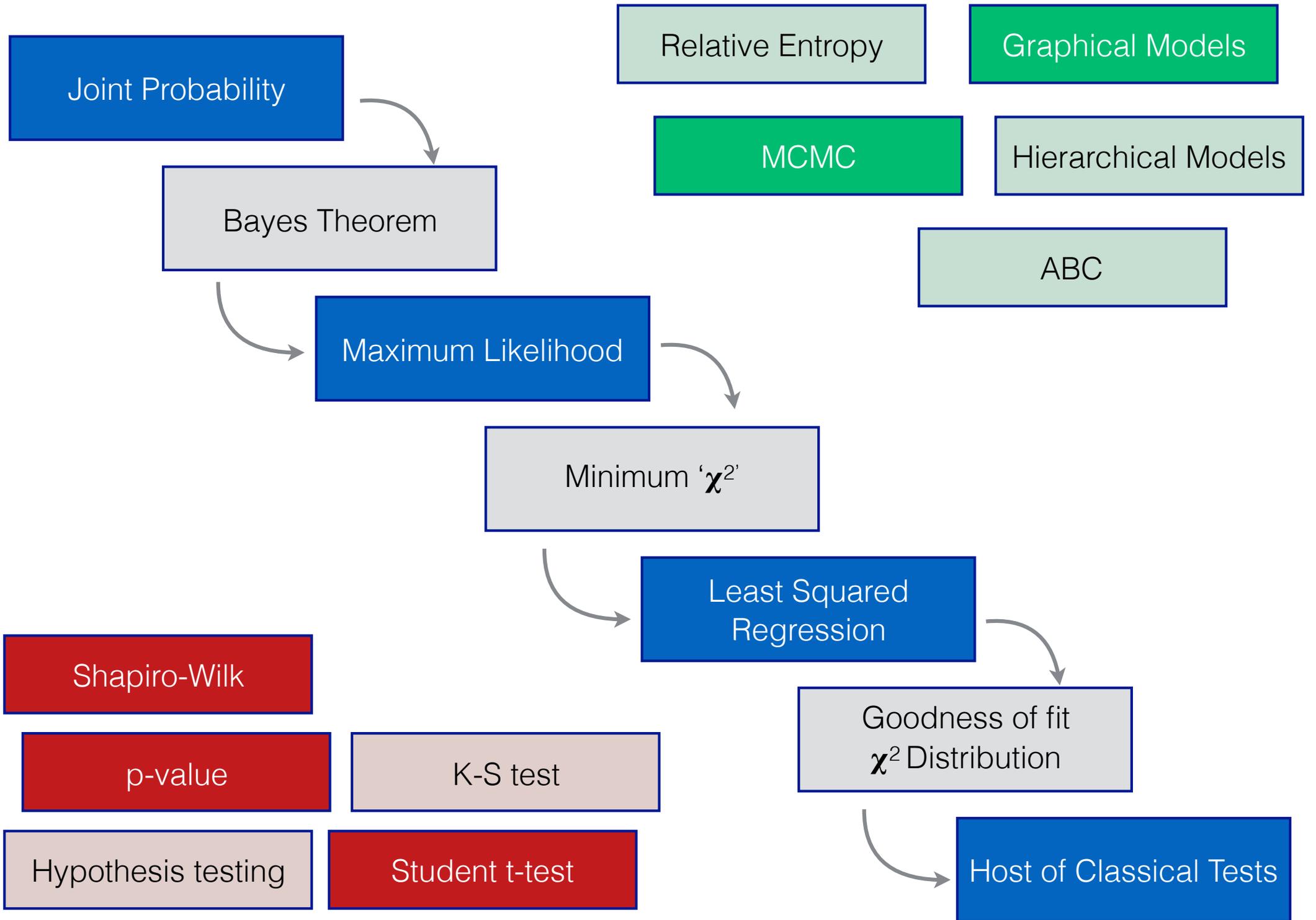
Student t-test

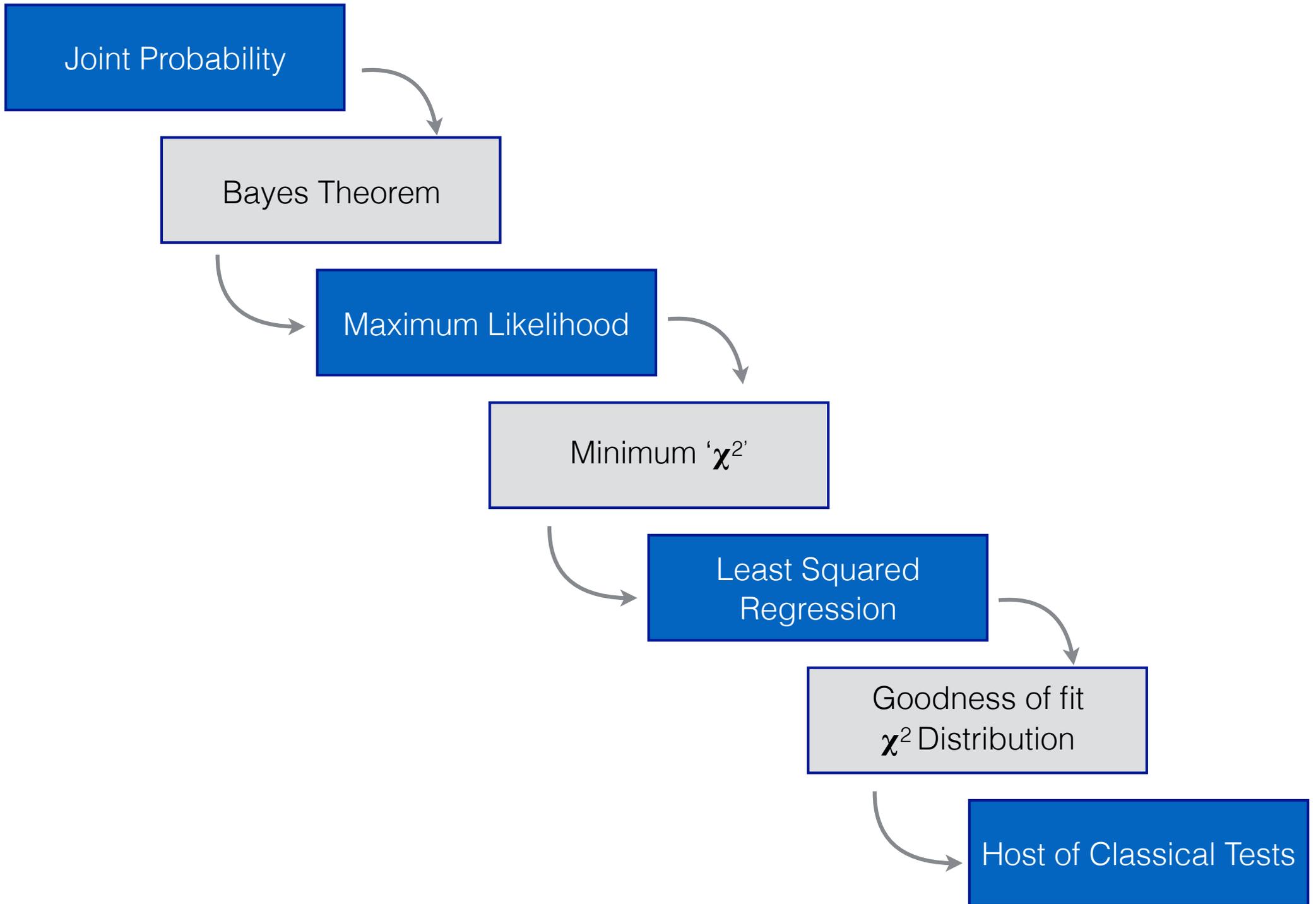
Shapiro-Wilk

Least Squared
Regression

ABC

Goodness of fit
 χ^2 Distribution





Probabilities

Joint

$$p(W, H)$$

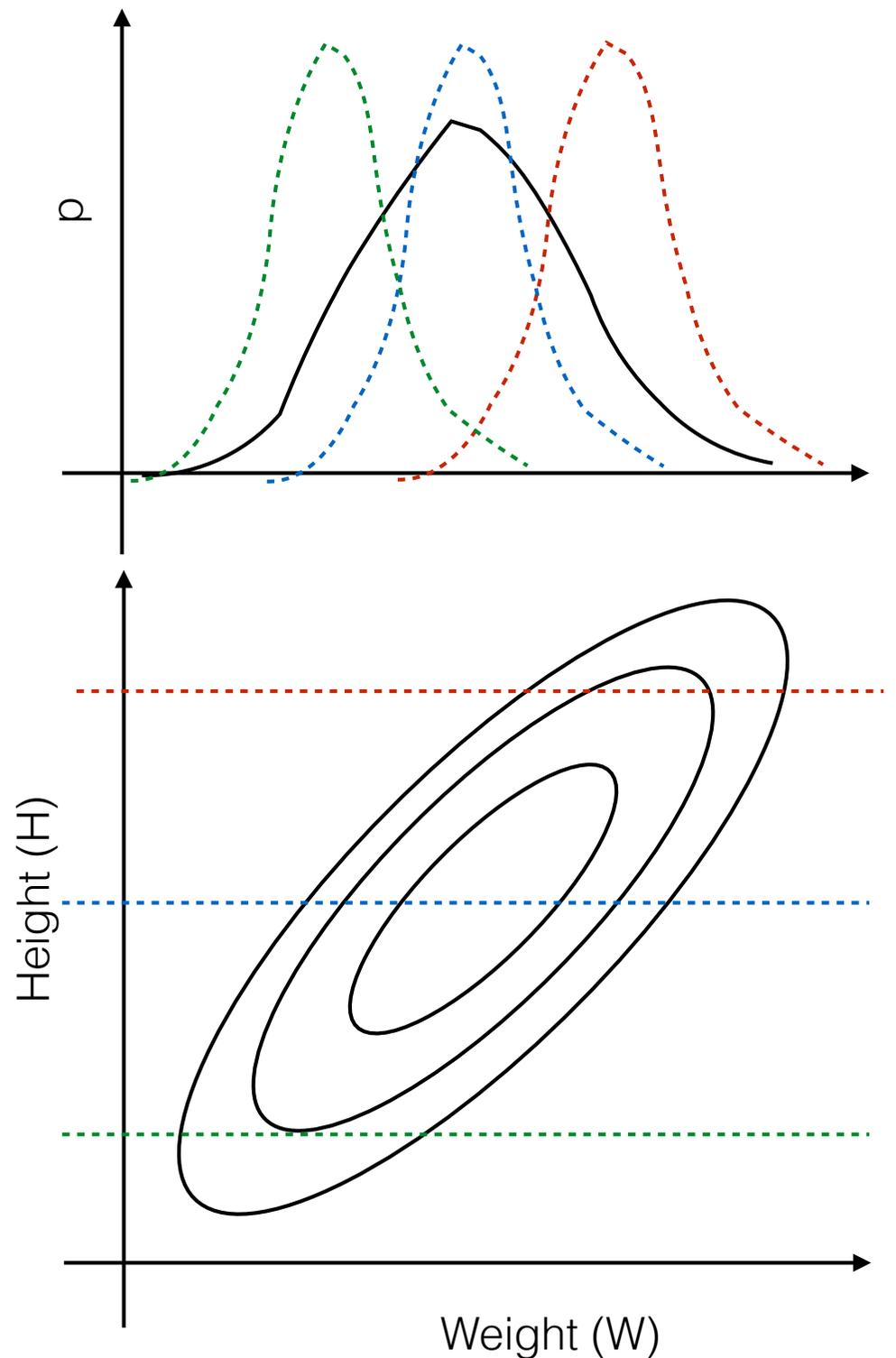
Marginalised

$$p(W) = \int p(W, H) dH$$

Conditioned

$$p(W|H)$$

$$\begin{aligned} p(W, H) &= p(H)p(W|H) \\ &= p(W)p(H|W) \end{aligned}$$



Bayes Theorem

Statistical inference problem:

Given some data (D) we wish to infer the parameters of a model (M)

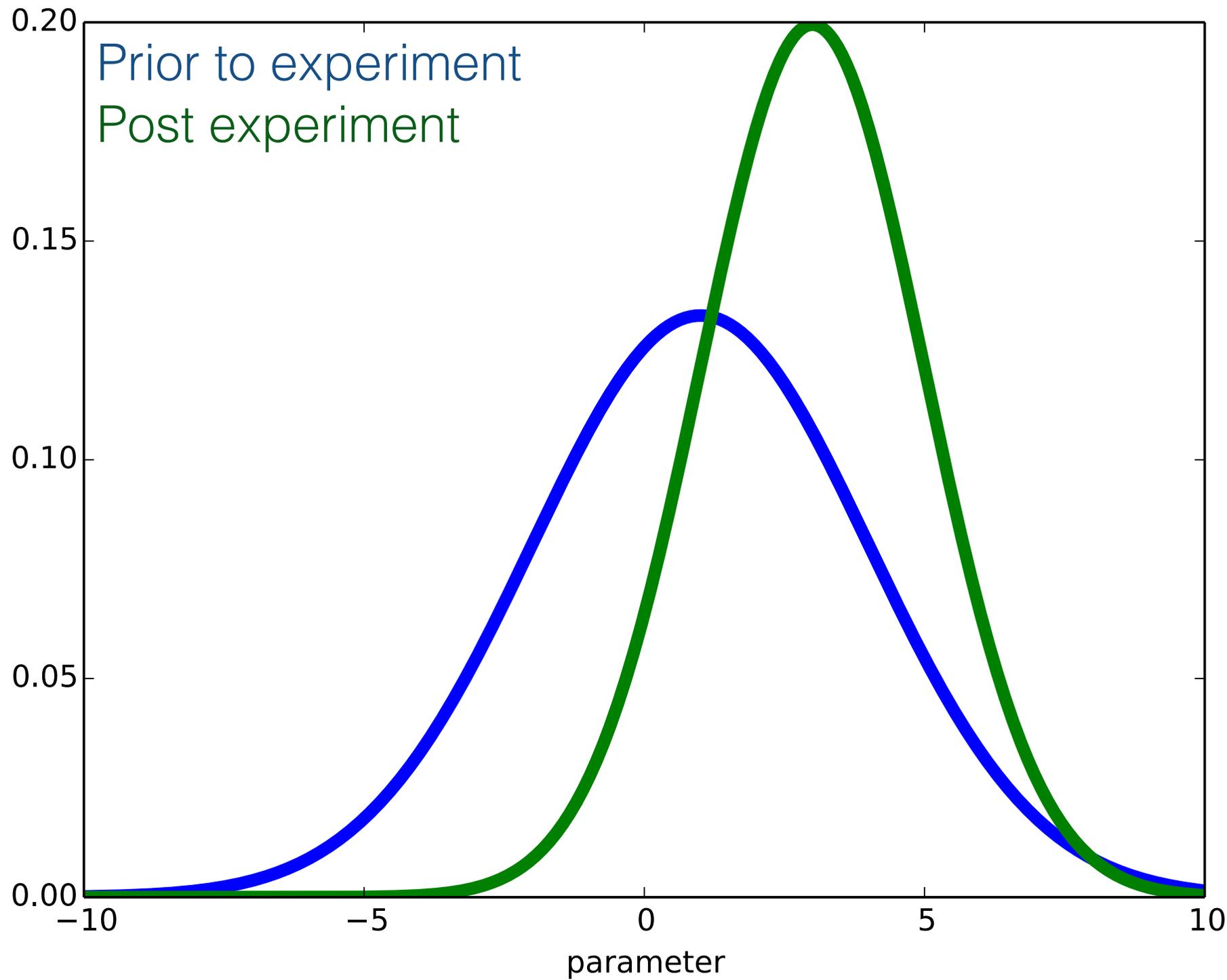
$$p(M, D) = p(M|D)p(D) = p(D|M)p(M)$$

$$p(M|D) \propto p(D|M)p(M)$$

Posterior

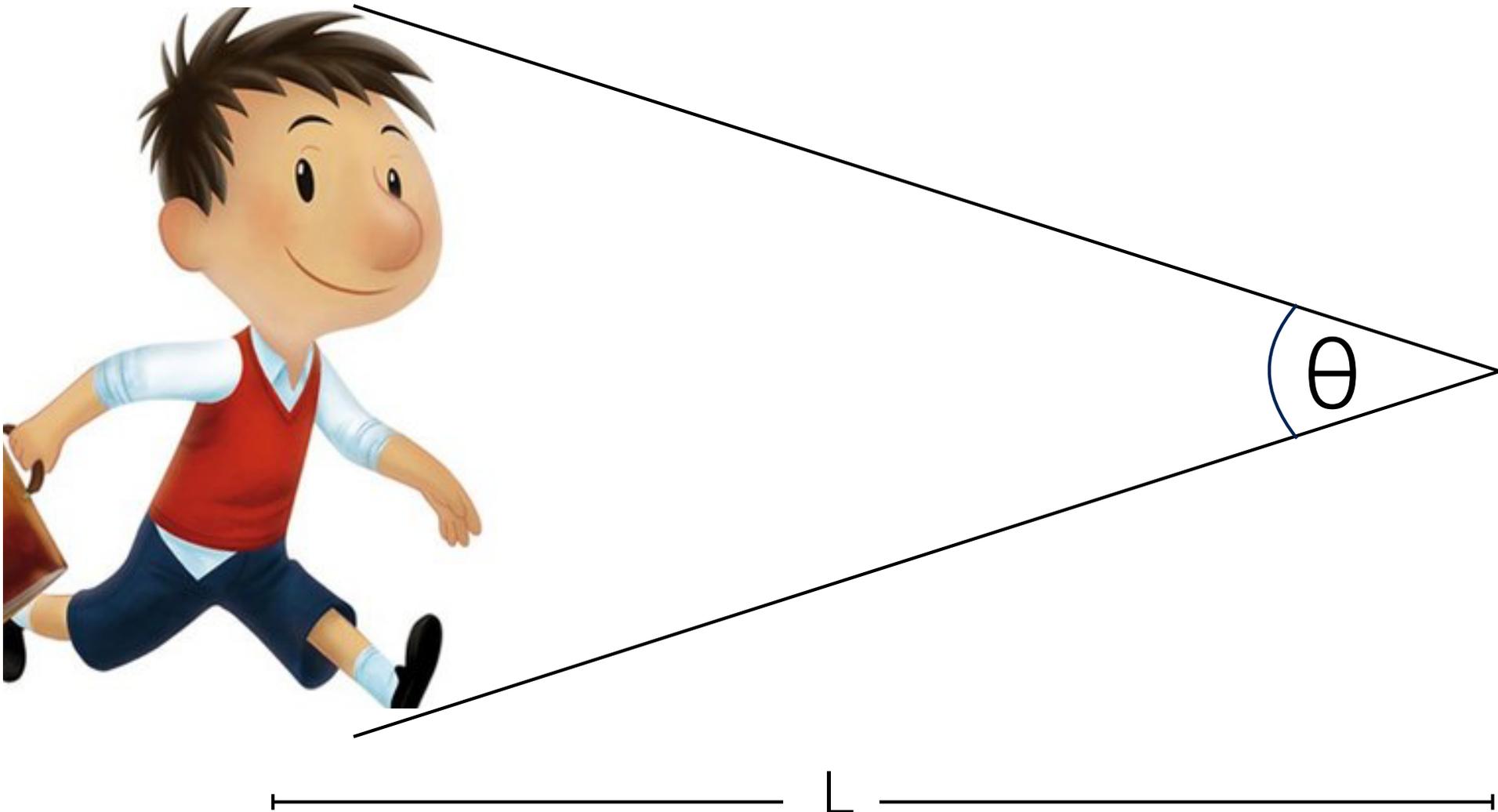
Likelihood

Prior

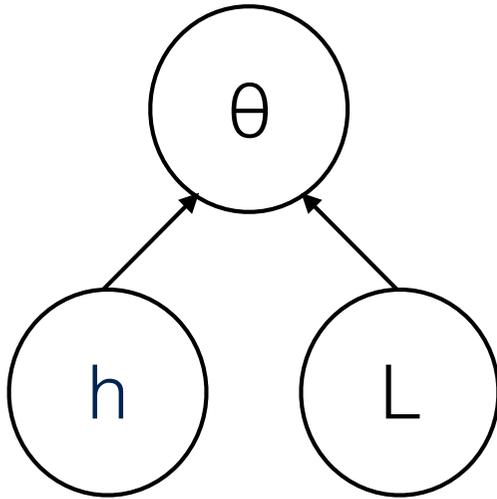


Petit Nicola

$$p(h|\theta) \propto p(\theta|h)p(h)$$

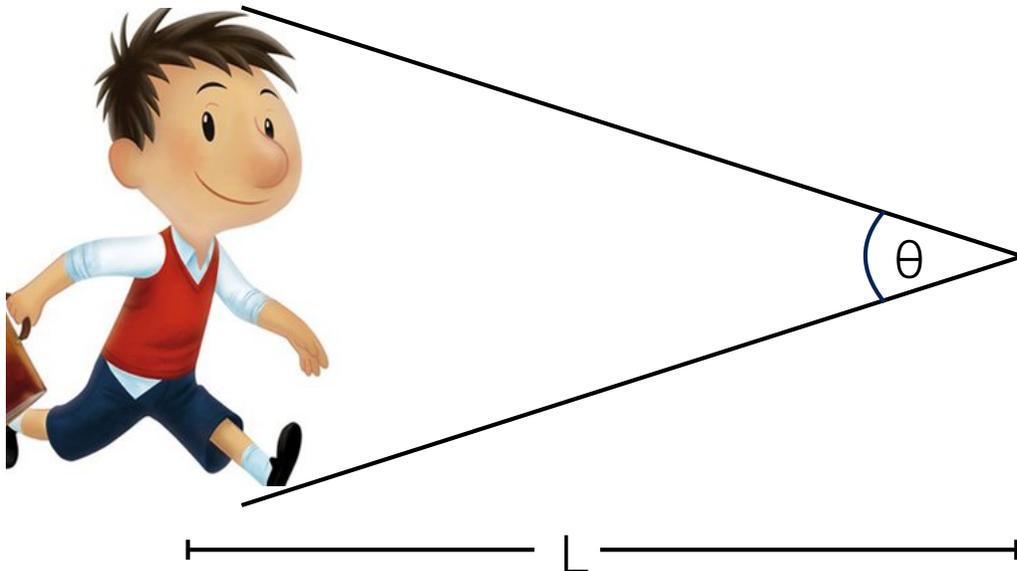


Petit Nicola



$$p(h|\theta) \propto p(\theta|h)p(h)$$

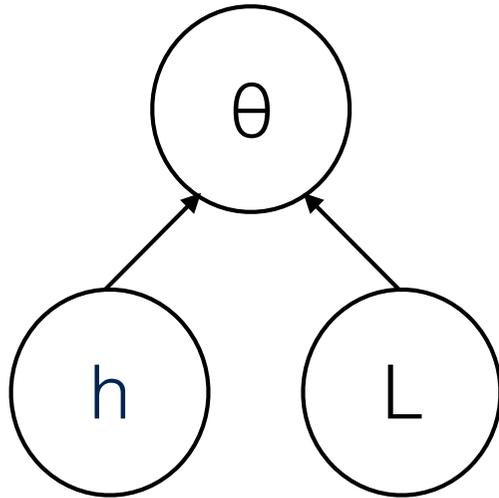
$$p(\theta|h) \propto \exp \left[-\frac{(h/L - \theta)^2}{2\sigma_\theta^2} \right]$$



$p(h)$ ← what do we know before making a measure

$$\left(p(h) = \int p(h, \theta) d\theta \right)$$

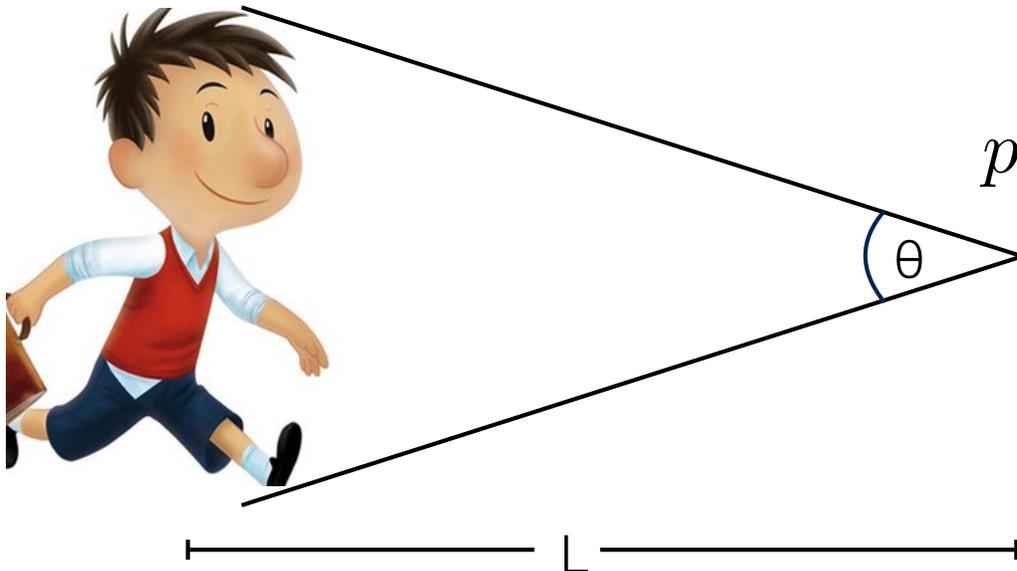
Petit Nicola



$$p(h|\theta)$$

$$p(h, L|\theta) \propto p(\theta|h, L)p(h)p(L)$$

$$p(h|\theta) = \int p(h, L|\theta)dL$$



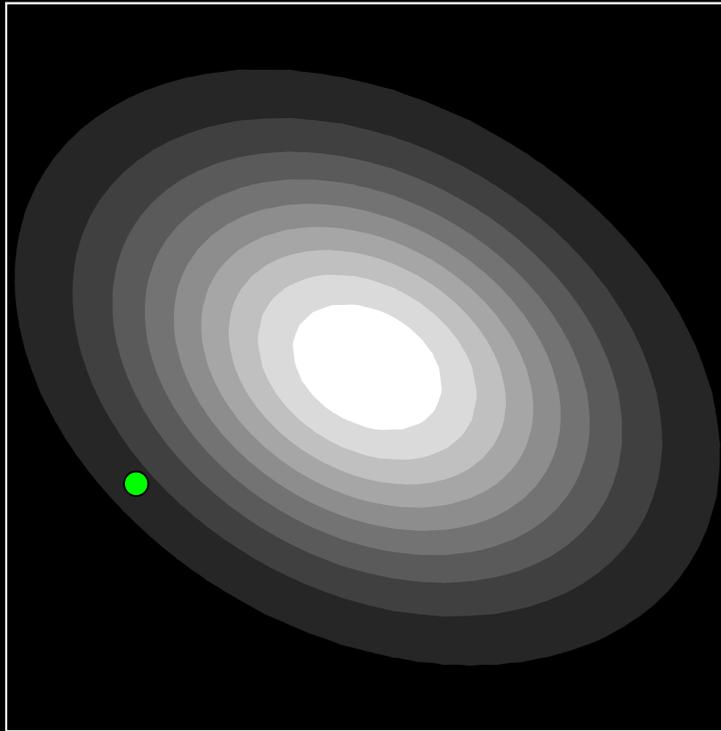
$$p(\theta|h, L) \propto \exp \left[- \frac{(h/L - \theta)^2}{2\sigma_\theta^2} \right]$$

MCMC - Point sampling of pdf

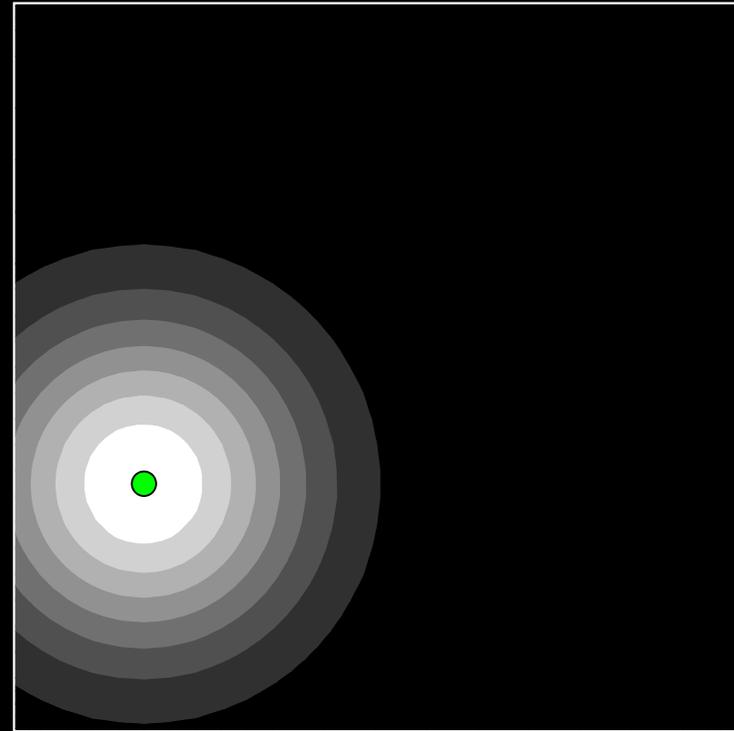
$$\langle x \rangle = \int x p(x) dx$$

$$\langle h(x) \rangle = \int h(x) p(x) dx$$

Metropolis-Hastings



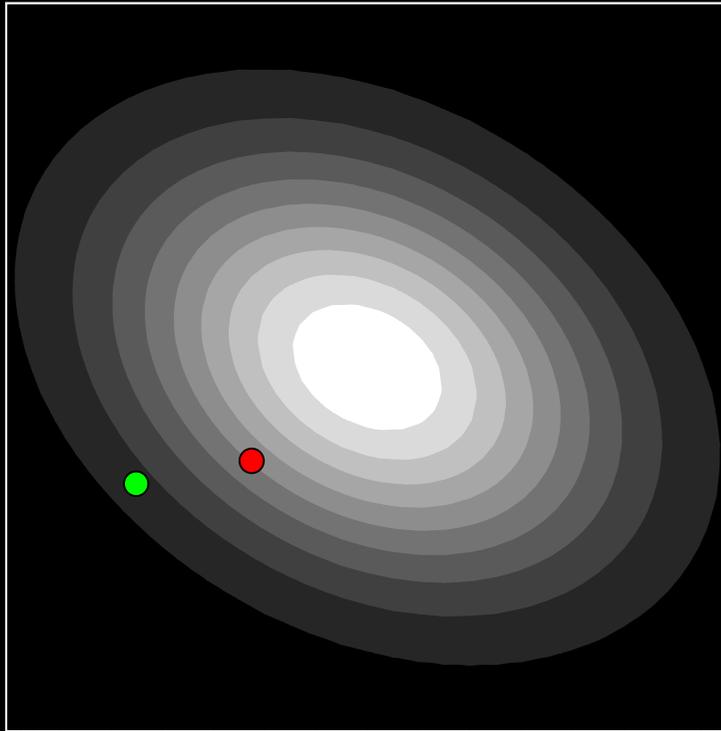
target distribution $P(\theta)$



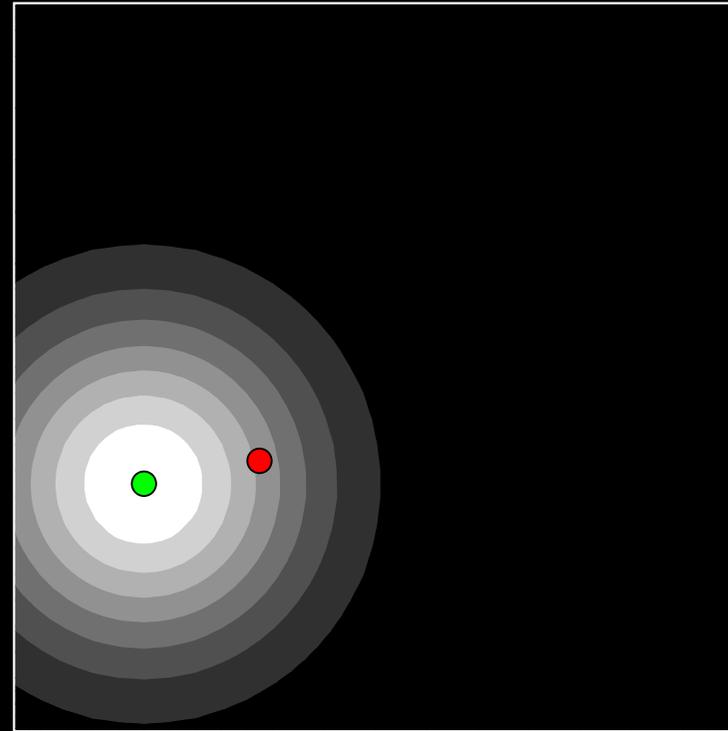
proposal distribution $Q(\theta)$

Metropolis-Hastings

initial position θ_0

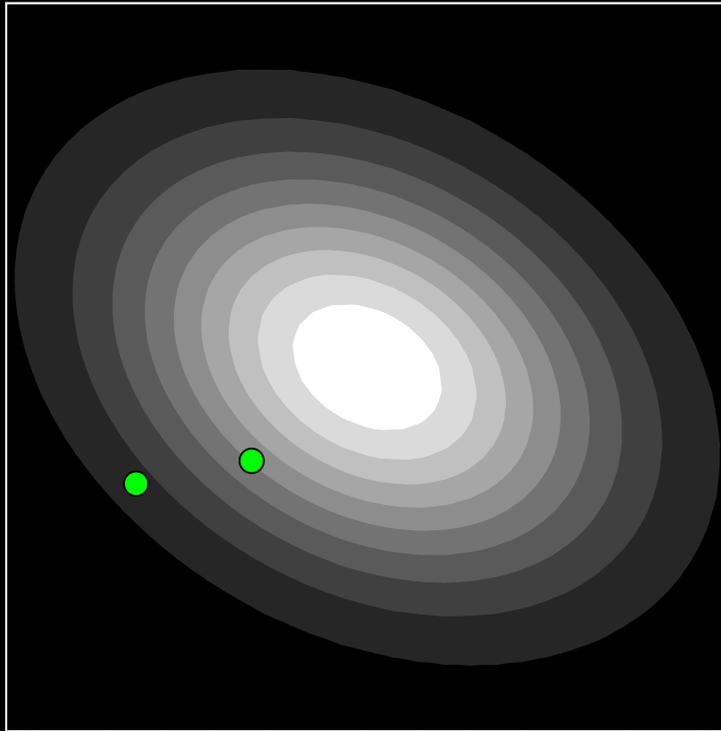


proposed position θ_p

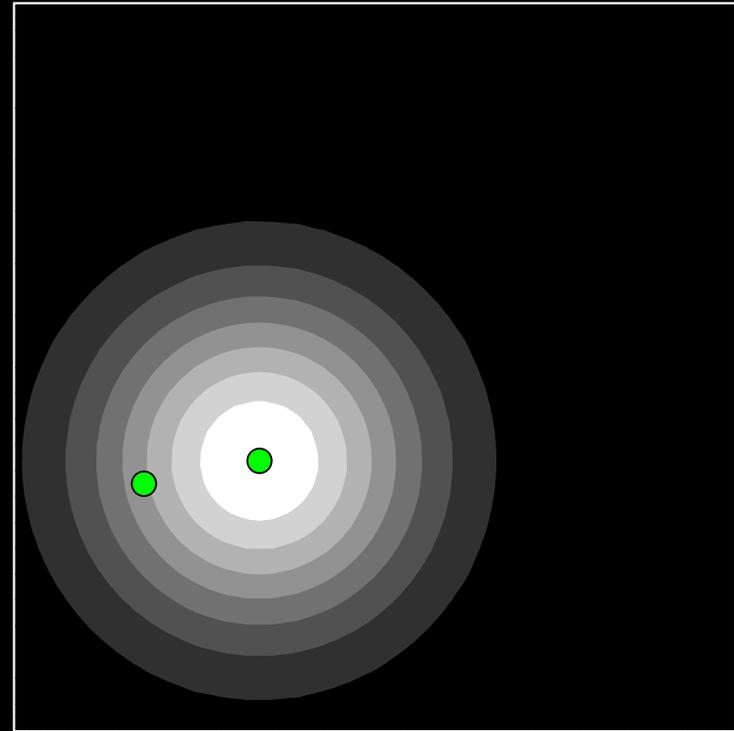


accepted with probability: $\min(1, P(\theta_p)/P(\theta_0))$

Metropolis-Hastings

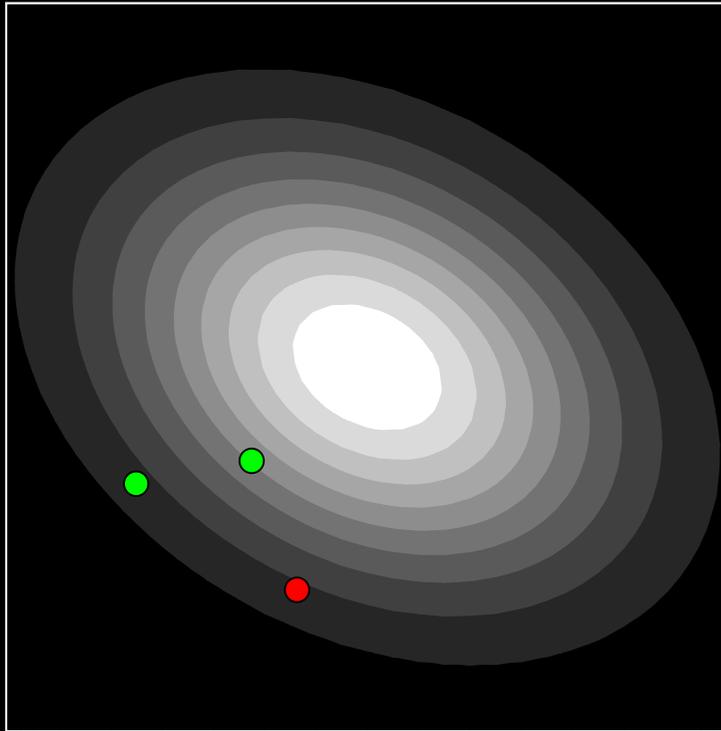


target

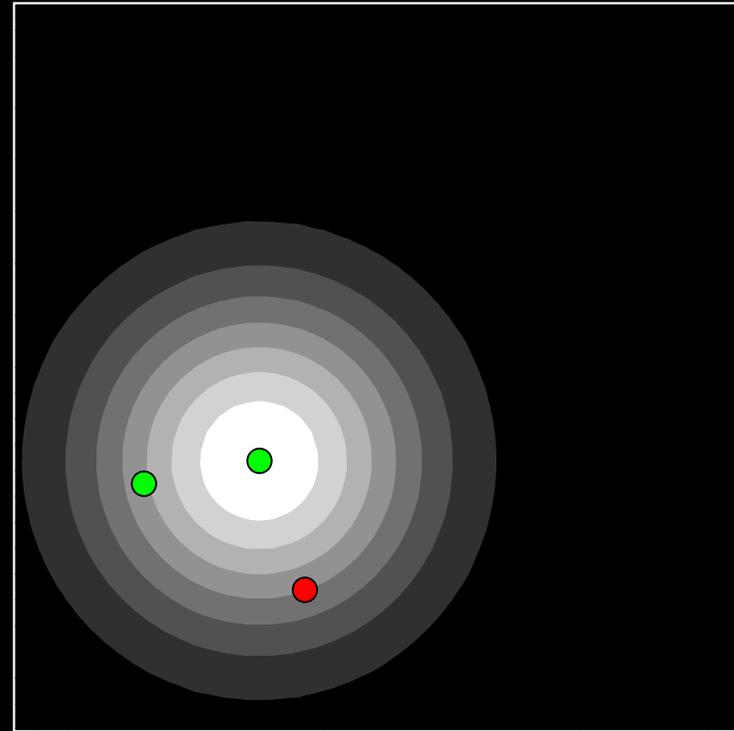


proposal

Metropolis-Hastings

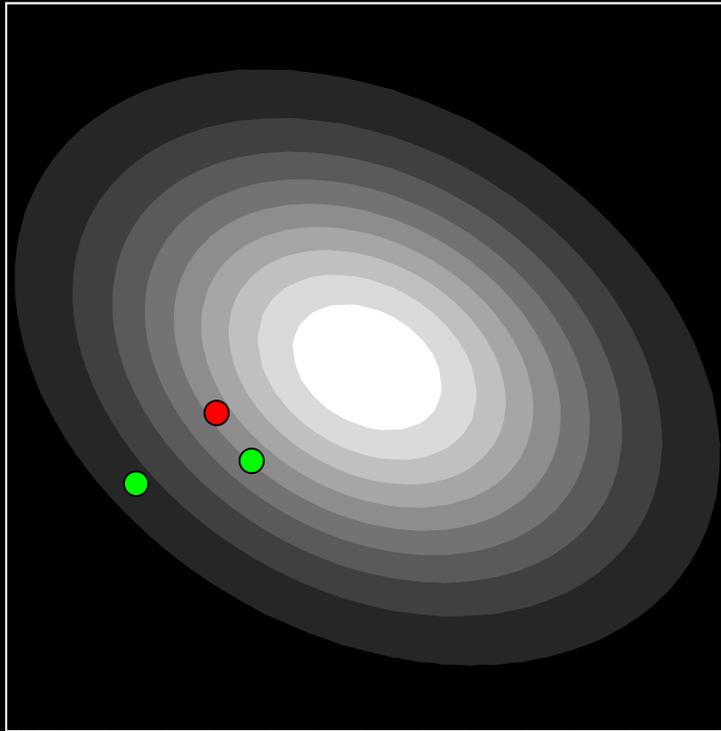


target

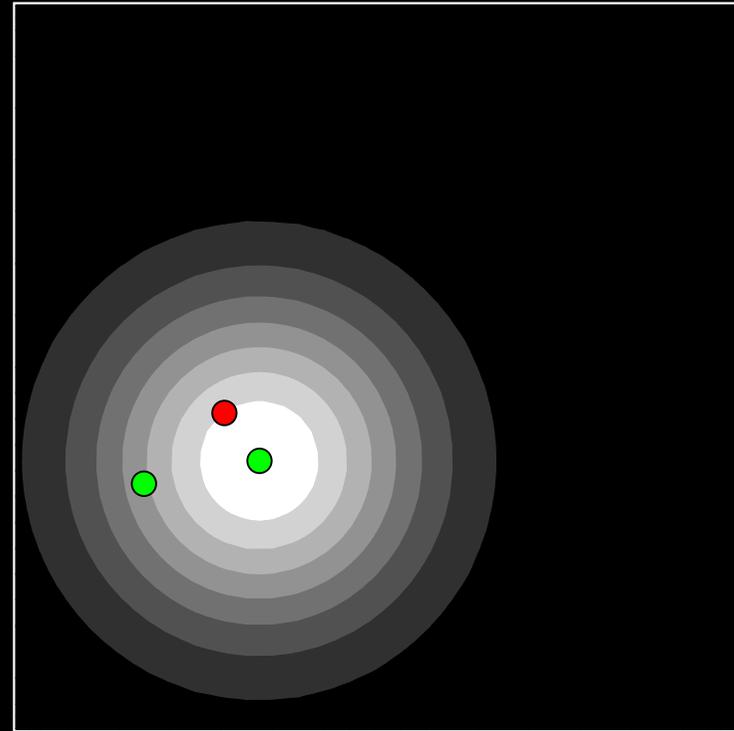


proposal

Metropolis-Hastings

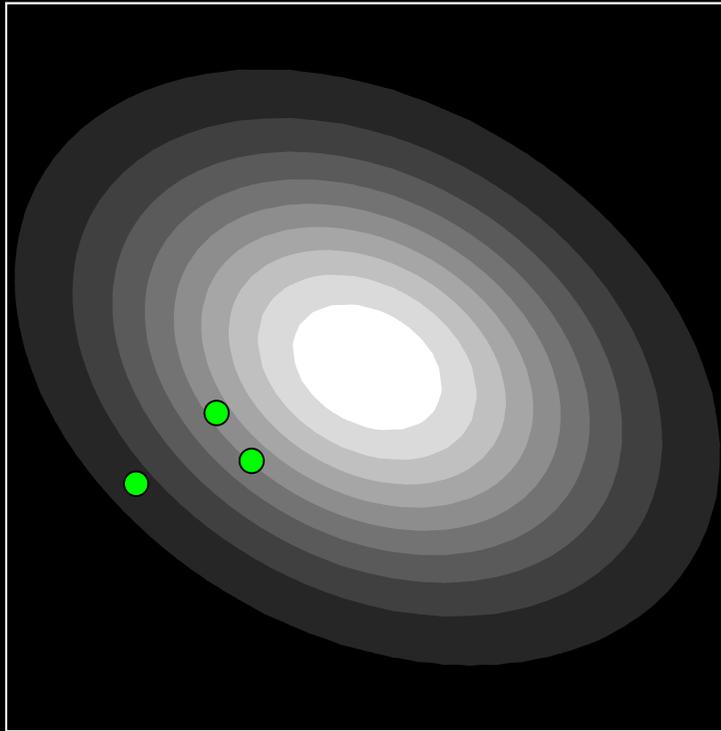


target

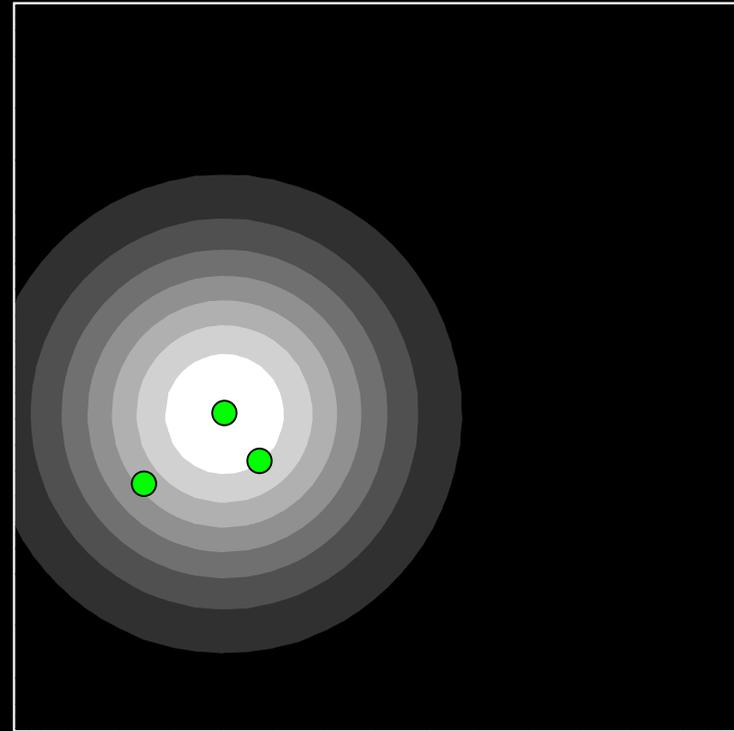


proposal

Metropolis-Hastings

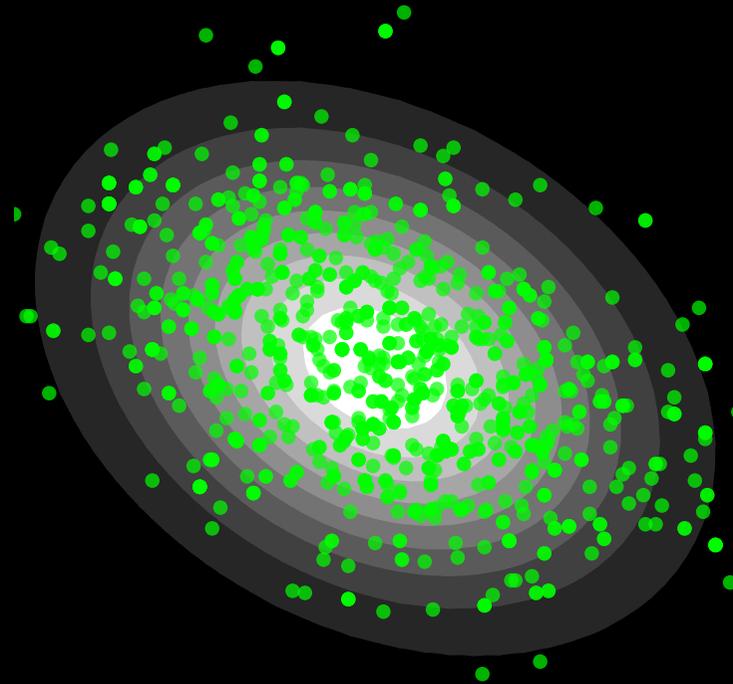


target



proposal

Metropolis-Hastings



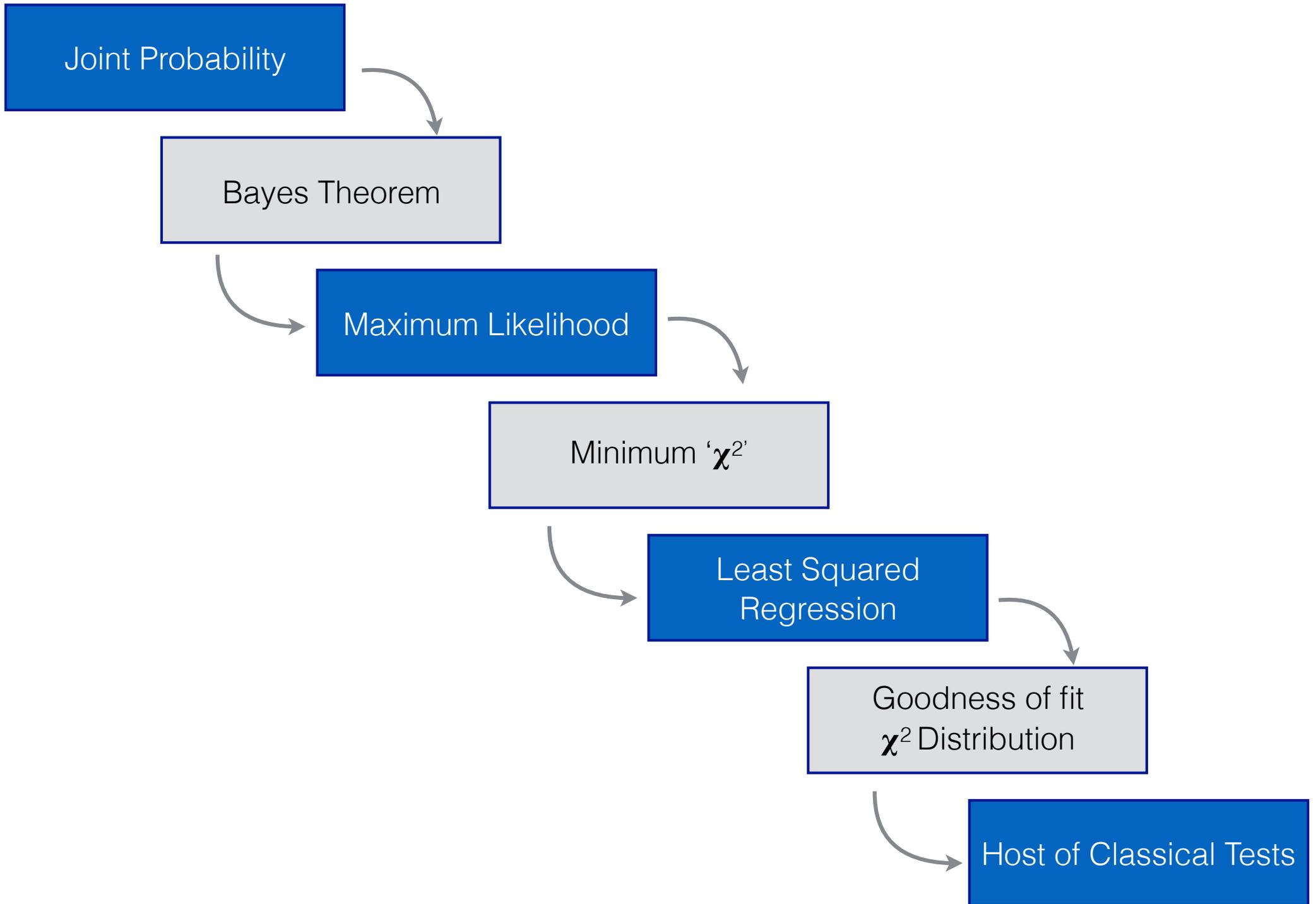
after 1,000 iterations

MCMC - Point sampling of pdf

$$\langle x \rangle = \int x p(x) dx$$

$$\langle h(x) \rangle = \int h(x) p(x) dx$$

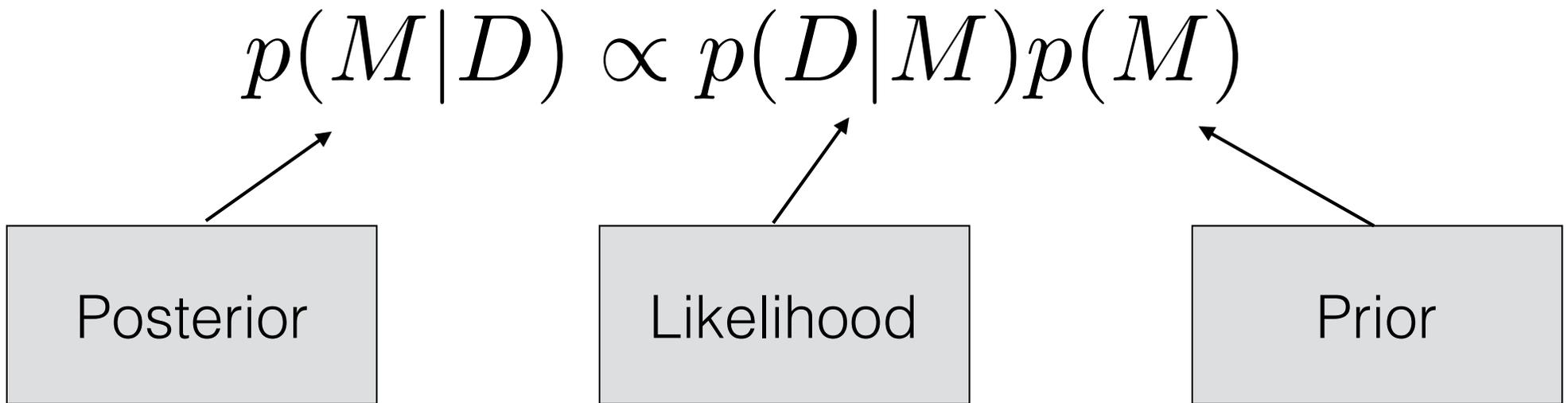
Other measures can also be extracted from the $p(x)$ e.g. x for max probability



Flat prior -> Likelihood analysis

$$p(M|D) \propto p(D|M)p(M)$$

Posterior



Likelihood

Prior

Maximum Likelihood

Posterior calculations with a flat prior and the maximum probably point recover a statistical method called the Maximum Likelihood Estimator (MLE) approach.

Though it was developed almost 100 years ago (R.A. Fisher 1920s), it is still a useful method for many applications.

Maximum Likelihood: IID process

$$L(\theta) = \prod f(X_i; \theta) = p(X_i | D)$$

$$\ell(\theta) = \ln[L(\theta)] = \sum \ln f(X_i; \theta)$$

The MLE of θ ($\hat{\theta}_{MLE}$) is found by maximising the $L(\theta)$ or $\ell(\theta)$.
 $\frac{\partial \ell}{\partial \theta} = 0$.

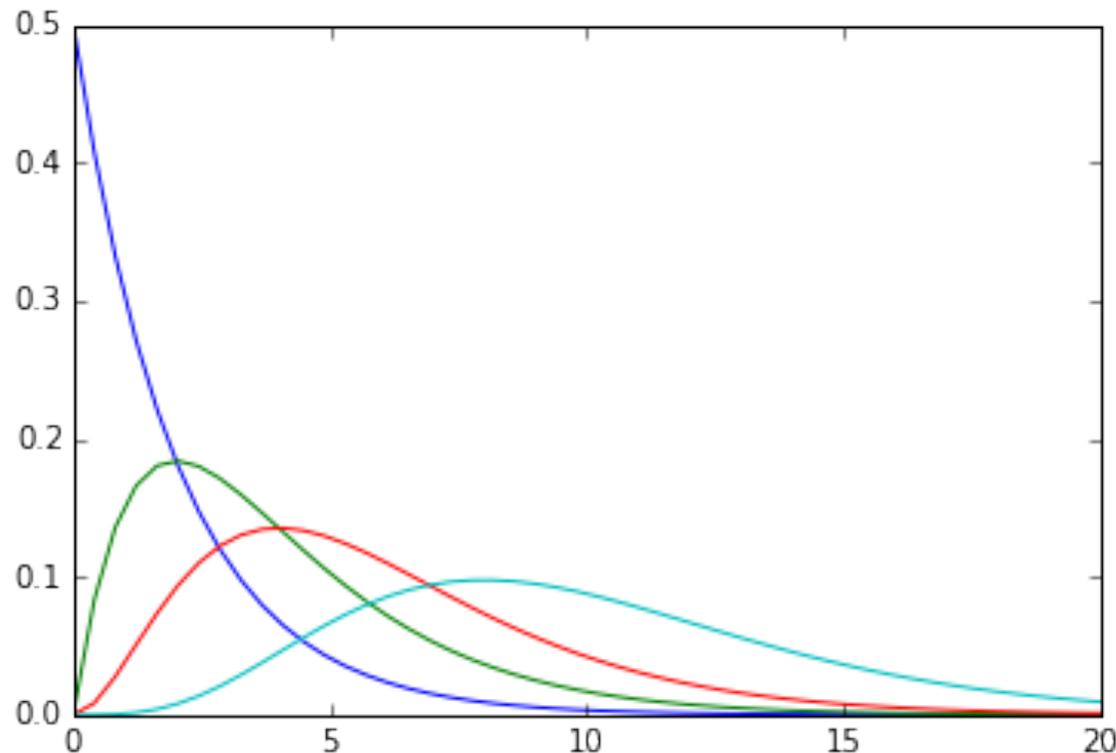
Maximum Likelihood: Gaussian example

Mean of a sample of size n from Gauss $N(\mu, \sigma^2)$

Minimum ' χ^2 ' & the χ^2 Distribution

Draw numbers from
a normal distribution:

$$l \propto - \sum \frac{(x_i - \mu)^2}{2\sigma^2} \quad \mathcal{N}(0, 1)$$



Interested in variable that
is the sum of squares:

$$Q_n = x_1^2 + x_2^2 + \dots + x_n^2$$

Distribution of Q_n follows a
 χ^2 distribution.

$$\frac{1}{2^{\frac{k}{2}} \Gamma\left(\frac{k}{2}\right)} x^{\frac{k}{2}-1} e^{-\frac{x}{2}}$$

Least Squared Regression - MLE for linear models

$$y = X\theta + \epsilon$$

$$\hat{y} = X\hat{\theta}$$

$$S_g = (y - X\hat{\theta})^T C^{-1} (y - X\hat{\theta}).$$

Where the known covariance, C , is

$$C_{ij} = \langle (y_i - \langle y_i \rangle)(y_j - \langle y_j \rangle) \rangle.$$

$$\frac{\partial S_g}{\partial \hat{\theta}} = 0$$

In this case the estimator is given by

$$\hat{\theta}_g = (X^T C^{-1} X)^{-1} X^T C^{-1} y,$$

with a covariance matrix of $(X^T C^{-1} X)^{-1}$

